Risk-averse Dealers in a Risk-free Market – The Role of Internal Risk Limits

Dan Li, Lubomir Petrasek, Mary Tian 1

March 2024

¹We thank Stephen Zoulalian and Karl Patram for excellent research assistance. We thank David Lynch, David McArthur and Dennis Mawhirter for explaining the details of the Volcker data to us. We thank David Bowman, Ellen Correia Golay, Norman Schürhoff and seminar participants at the Federal Reserve Board (Feb. 2024) and University of Lausanne (March 2024) for their comments. The views of this paper do not represent the views of the Federal Reserve Board.

Abstract

Self-imposed risk limits effectively limit dealers' appetite for risks and their capacity to intermediate in Treasury markets in times of market stress. Using granular and high frequency regulatory data on US dealers' Treasury securities trading desk positions and desk-level Value-at-Risk limits, we show that dealers are more inclined to reduce their positions as they get closer to their internal risk limit, consistent with such limit being meaningful and costly for traders to breach. Dealers actively manage their inventories away from their limits by selling longer-term securities and requiring higher compensation to take on additional risks. During the height of the Covid-crisis in 2020, dealer desks that were closer to their VaR limits sold more Treasury securities to the Fed and accepted lower prices in the emergency open market operations. Our findings complement studies that link post-GFC bank regulations to market liquidity by showing that self-imposed risk limits can explain the risk-averse behavior by dealers, and offers a micro-foundation for the link between market volatility and market liquidity in dealer-intermediated OTC markets. Policy prescriptions such as deregulation alone may not be sufficient to induce risk-taking by dealer intermediaries in times of crisis.

JEL Classification: G01, G23, E52

Keywords: Dealer Intermediation Capacity, Treasury Market, Risk Limits, Regulation, Market Liquidity

1 Introduction

In March 2020, the US Treasury market, long heralded as the paragon of liquidity and the a haven of safety, experienced surprising turbulence when foreign central banks and institutional investors such as mutual funds rushed to liquidate their most liquid assets in the face of the global pandemic (Vissing-Jorgensen, 2021). Prices and liquidity of Treasury securities declined sharply as one-sided client order flow met hesitant dealers. Even primary dealers, committed market makers of these securities, exhibited an increased reluctance to "lean against the wind" and buy Treasuries. Indeed, as shown in Figure 1, dealers sought to reduce their inventory of Treasury securities during the crisis even as Treasury yields rose, contributing to the selling pressure.

Many attribute this perplexing apathy of dealers to the slew of post-Global Financial Crisis (GFC) bank regulations, prominently the Volcker rule and the Supplementary Leverage Ratio (SLR). However, the Volcker rule that prohibits dealers from proprietary trading explicitly exempts US Treasury securities, in addition to exempting market-making activities. An impact on dealers' Treasury market making activities of enhanced capital regulations, especially the U.S. SLR, seems more plausible. Indeed, the SLR is "risk blind"— it treats traditional risk-free assets like Treasury securities similarly to their riskier corporate debt counterparts, which makes intermediation in low-margin high-volume risk-free market particularly costly (Duffie, 2018).

Yet, this narrative remains incomplete. The SLR itself is slow-moving, with the on-balance sheet exposure being calculated as the average of the the daily values over the entire quarter. A temporary increase in dealer inventories has only a modest effect on the quarterly regulatory ratio. The SLR is also insensitive to risk—a sudden rise in market volatility does not alter SLR. Some argue that sharp deposit inflow at the commercial bank siblings of broker-dealers during the onset of the Covid crisis inflated banks' balance sheet, making their SLR more binding. Dealer arms of the same bank holding companies may have reacted by limiting their market-making activities (Breckenfelder and Ivashina, 2021). However, there was substantial headroom in bank SLR in March 2020 and the headroom was larger then than in the subsequent years, when the Treasury market was relatively calm. Moreover, high frequency evidence in Vissing-Jorgensen (2021) points to the timing of a reversal of price pressure in Treasury market in 2020 that lines up well with a large increase in the Federal Reserve's daily Treasury purchases on March 19, weeks before regulatory reliefs that temporarily exempted Treasuries and reserves from the calculation of SLR for banks were announced.

Moreover, the reluctance of dealers to embrace risk in times of crisis, even amidst the haven of the most risk-free assets, predates regulatory oversight. Recent stress events in the Treasury markets show a remarkable resemblance with historical events. For example, in September 1939, dealers proved unable to intermediate the surge in sales of Treasury securities at the onset of World War II, requiring the Federal Reserve to conduct large scale purchases of Treasury securities to preserve orderly market conditions and alleviate pressures at dealers (Duffie, 2023).

Our study reveals that dealers are inherently risk averse — they employ internal risk limits, which effectively constrain their risk absorption capacity. These limits, typically imposed by firms' risk management function, are carefully set at granular trading-desk level, and portfolio risks relative to these limits are scrutinized daily. We show that internal risk limits are persistent—that they do not change often—although they are not completely rigid. The limits sometimes get relaxed, especially during market stress event, but such relaxation of the limits is frequently insufficient to fully offset the demands on dealer risk taking. Empirically, we show that such limits are not toothless—traders are mindful of the various limits and triggers and tend to manage their portfolio in ways that avoid breaching such limits. Such evidence indicates that breaching the limit carries substantial costs to traders, the costs could be either reputational or pecuniary.

The schemes of internal risk limits that trading desks are subject to could be quite elaborate, with potential limits on market risks, counterparty risks and liquidity risks; the actual schemes vary substantially across firms and business lines. For example, derivatives trading desks tend to have limits with regards to various portfolio greeks (delta, vega, gammas). Forex trading desks might have limits with respect to foreign currency exposures. We focus on a particular risk limit that is universally adopted by almost all Treasury securities trading desks — the Value-at-Risk (VaR) limit. In addition to being universally adopted by all desks, VaR is also considered as the primary risk limit that traders pay attention to. By design, VaR captures the magnitude of potential tail losses, and is highly sensitive to sharp changes in market volatility and is intrinsically counter-cyclical. A static portfolio would have substantially higher VaR should market volatility rise. Our unique data on desk-level risk limits have a few advantages. By observing the upper bound on VaR and its actual usage at the trading desks level, we can measure dealer intermediation capacity directly, rather than inferring it. The high frequency panel nature of the data also allow us to empirically identify effect of risk limit on risk-taking behavior better than looking at market aggregate.

We show that dealers avoid breaching desk VaR limits by managing down their inventory away from the upper limit, and in particular, by acquiring less inventory or selling existing inventory, with newer inventory being sold first. Our VaR-based result provides a micro-foundation to the negative link between market volatility and dealer intermediation capacity.

Notably, since VaR calculations reflect portfolio risk after taking into account offsetting positions and hedges, one would expect the net positions of dealers to be more important to manage than gross positions, in order to keep their VaR in check. Our findings confirm that dealers' net positions are indeed more affected by their VaR constraint than their gross positions. A size-based constraint that does not net long and short positions, such as the SLR constraint, would have the opposite implication for dealer behavior, with gross positions being more important to manage than net positions. These findings also suggest that there might be bounds in what policy recommendations that are aimed at alleviating dealer balance sheet pressures could achieve during market stress. For example, mandatory central clearing or a relaxation of the minimum SLR could provide relief from size-based regulatory constraints without alleviating risk-based constraints.

To further demonstrate that large Treasury market dealers faced significant constraints due to their internal VaR limits in March 2020, we conduct an event study. We compare how trading desks were differently inclined to reduce their inventories through selling to the Federal Reserve during the emergency open market operations from March to June. Our analysis shows that dealers closer to their VaR limits were more inclined to sell to the Fed in the subsequent days and months. These dealers also offloaded larger amounts of Treasury securities to the Fed's balance sheets, both in absolute quantity and relative to their overall intermediation activities. The sensitivity of dealers' sales to the Fed relative to the tightness of their VaR constraint was stronger in March, than from April to June.

In addition to selling more to the Fed, more constrained dealers also submitted more aggressive

offers in System Open Market Account (SOMA) auctions — they were willing to sell their Treasury bonds to the Fed at cheaper prices. We find that the aggressiveness of dealers' auction offers was particularly sensitive to their VaR limit usage for longer-maturity bonds. These bonds contribute disproportionately to the VaR of the portfolio and generate the most relief from VaR limits should they be disposed of from dealers' portfolios. Such evidence of cross-sectional differences further points to VaR limits, instead of risk-blind size-based constraints, as being the more relevant constraint during the worst period of market illiquidity in 2020.

How do dealers' VaR limit and their trading with the Federal Reserve affect dealers' provision of liquidity to clients? We first confirm that during the Covid-crisis period, dealers were less likely to provide liquidity to clients (buy from clients) if they were closer to their VaR limit. However, after controlling for VaR limit usage, dealers who sold more to the Fed were more likely to purchase from clients, especially in March, 2020. Therefore, by alleviating the balance sheet constraints of dealers and taking bonds off dealers' inventory, the SOMA purchase operations were effective in incentivizing dealers to provide more liquidity to clients. Interestingly, dealers who were able to sell more bonds to the Fed were able to provide more immediacy to their clients for Treasury securities both in and outside the SOMA operations. This suggests that dealers were not merely passing bonds from clients to the Fed. Since dealer balance sheets are fungible, Fed support in some bonds has positive spill-over effects on other bonds — the extra capacity vacated from selling some bonds to the Fed allowed dealers to provide more immediacy to clients in other bonds.

We then explore how risk aversion among dealers translates to customers' transaction costs. We show that as dealers approach their VaR limit, they seek greater compensation for taking on any additional risk. We show that dealer average future profit and loss (P&L) scaled by their desk VaR, increases with the tightness of their VaR constraints. Moreover, the effect is entirely driven by P&L generated from new trades, rather than existing positions. Since the flip side of dealer profitability from new trades is client transaction cost, this result offers a direct link between dealer capacity constraint and transaction costs.

Finally, we develop an aggregate measure of Treasury market dealer VaR constraints using individual dealer VaR limit tightness and its recent share of customer transactions as weights. This aggregate measure exhibits a strong correlation with established measures of Treasury market illiquidity for both average and tail measures of illiquidity, even after market volatility is controlled for. Furthermore, consistent with Duffie, Fleming, Keane, Nelson, Shachar, and Van Tassel (2023), we find that the impact of our aggregate dealer VaR constraint measure on Treasury market illiquidity is skewed and more pronounced if market illiquidity is at the higher quantiles.

2 Related Literature

Our findings challenge the classic micro-structure assumption that dealers are risk neutral. We argue that the apparent "risk averse" behavior of dealers in time of stress, even in "risk free" asset market, is consistent with them optimizing under constraints. Duffie (2023) extends the dealership model of Amihud and Mendelson (1980) to include an explicit constraint on dealer inventory, where the dealer's objective is the maximization of the present value of market making profits under the constraint that the inventory should be strictly below a threshold. Duffie (2023) shows that the simple extension allows the model to generate dealer bid-ask quotes that are significantly wider when dealer inventory is near capacity.

We contribute to the vast literature of OTC market intermediary behavior by offering a peek behind the curtain of the real world of dealer trading desks and internally-set risk limits. Such risk limits allow us to measure the spare capacity of dealers directly, rather than inferring it. We document some stylized facts about desk-level VaR limits, how persistent they are, their general level of usage and how it varies across time and across desks. Anderson, McArthur, and Wang (2023) look at several risk limits used by corporate bond trading desks and show that they affect corporate bond spreads and transaction costs.

Motivated by the sharp price movements and surprising hiccups in Treasury market functioning in March 2020, we focus on Treasury securities trading desks of major dealers and attempt to gain a deeper understanding of the choices they make and the constraints they face in different market environments and how dealer decisions translate into market outcomes such as market liquidity. Our study complements the literature that studies the behavior of various institutional investors of Treasury securities during the March 2020 market stress. Falato, Goldstein, and Hortaçsu (2021) and Kruttli, Monin, Petrasek, and Watugala (2021) document large sales of Treasury securities by mutual funds and hedge funds, respectively. Using data on both the price and quantity of liquidity in the Treasury market, Goldberg (2020) is able to decompose liquidity into supply and demand. He estimates that investors' demand for liquidity soared, while dealer liquidity supply declined notably during the on-set of the pandemic in March 2020. We provide micro-level evidence that attributes the lack of dealer liquidity provision to their internal risk limit.

In addition, our paper is related to the growing literature on intermediary balance sheet constraint and market liquidity. He, Nagel, and Song (2021) highlights the importance of intermediaries balance sheet costs in explaining the disruptions in March 2020 in the Treasuries market. In particular, they point to the SLR as a potential source of dealer balance sheet constraint. Breckenfelder and Ivashina (2021) exploit the persistence in client-dealer relationship associated with home country bias and show that European corporate bond mutual funds with closer ties to banks with constrained leverage ratios experienced more outflow and selling pressure. Our research provides evidence that trading desk risk limits present another source of significant dealer constraint and documents how dealers manage their trading portfolios and set their required returns to avoid breaching such limits. We stress that the internal risk constraint and the regulatory constraint are not mutually exclusive. Both types of constraint could be at work simultaneously. Indeed, Duffie, Fleming, Keane, Nelson, Shachar, and Van Tassel (2023) show that a variety of aggregate dealer capacity measures can explain Treasury market liquidity beyond what can be explained by market volatility. Instead of looking at market aggregates, we focus on the cross-section of dealers and individual trading desks, and attempt to establish micro-level links between capacity constraints due to internal VaR limits and dealers' provision of liquidity.

Our findings are also consistent with the literature documenting the effectiveness of Federal Reserve open market purchases in alleviating Treasury market stress. Indeed, by carefully tracing out the timeline of Treasury securities price movement during the COVID crisis and the announcement and actual implementation of various policy actions, Vissing-Jorgensen (2021) identified a causal link between the Federal Reserve's Treasury purchases and market stabilization, and ruled out other potential confounding factors. We contribute significantly to this literature by linking dealer behavior, internal risk limit and SOMA auctions at the most granular level. We show that dealers more constrained by the VaR limit are more likely to sell bonds to the Fed and offer more

aggressively in auctions. Those who were able to sell more to the Fed were able to offer more immediacy to clients. Thus, the SOMA purchase program was highly effective in stabilizing the market by providing direct relief of dealers' balance sheet constraints.

This paper provides micro-level evidence that complements Adrian and Shin (2013), who document a negative relationship between banks' leverage and VaR, and that banks deleverage substantially during market downturns. Their findings, while focusing on slower-moving lending by banks through business cycles, is consistent with what we find for faster-moving activities at trading desks. The authors empirically document that banks keep the ratio of their VaR- to-equity roughly constant and develop a contracting model that allow a bank creditor to impose a VaR limit that is proportional to the banks' equity. Such contracting arrangement alleviates the risk-shifting moral hazard of the bank, and can effectively bound the bank's probability of failure below a fixed threshold, irrespective of the risk environment. This again mirrors our finding that trading desks not only universally adopt a VaR-type risk limit, but they also take the internally imposed limit seriously and rarely change it.

While we do not model or attempt to explain why dealer banks impose strict risk limits on their trading desks, agency problems between traders and bank shareholders or creditors resemble the contracting frictions addressed in Adrian and Shin (2013). Traders are typically awarded a proportion of the trade profits, but have limited liability and are not expected to absorb large losses. Such compensation schemes potentially lead to excessive risk taking. Desk-level VaR-type limit is a natural way for dealer firms to alleviate the agency problem commonly discussed in the corporate finance literature. However, an ex-ante commitment to risk limits may come at the expense of forgoing profitable trading opportunities during market downturns. Similar types of risk management practices exist for other settings of delegated portfolio management. For example, large hedge funds typically have VaR-type risk limits for their traders as well, and while some hedge fund or high frequency traders may provide liquidity in normal market conditions, it is well documented that they typically retreat when market volatility rises (Boyson, Stahel, and Stulz, 2010; Aragon and Strahan, 2012; Kruttli, Monin, Petrasek, and Watugala, 2021). Agency problems within the corporate structure of asset managers or dealer firms induce procyclical risk appetite and renders such entities unsuitable to serve as "buyers of last resort" (Shleifer and Vishny, 1997). The rest of this paper is organized as follows. Section 3 describes the data. Section 4 shows main results about dealer inventory and dealer VaR usage. Section 5 presents an event study of dealer behavior after Fed purchases in March 2020. Section 6 examines the effects of internal risk limits on dealer profitability and aggregate market liquidity. Section 7 concludes.

3 Data

3.1 Trading Desk Risk Limit Data

Our main data comes from Volcker Rule collection. To insulate banks from the volatility and potential large losses originated from capital markets activities, the Dodd–Frank Act includes a prohibition known as the "Volcker Rule" on proprietary trading by U.S. banks. The rule allows exemptions for hedging, market-making, and various financial instruments such as foreign exchange and government securities. To facilitate the monitoring of compliance with the Volcker rule, dealers affiliated with large bank-holding companies are required to report daily metrics for each of their trading desks on Form FR VV-1. This report collects detailed data from trading desks, including a daily time series of various risk limits and their usage, profit & loss and its attribution to risk factors, and trading positions and transactions.

3.1.1 Treasury Market-making desks

While Treasury securities are exempted from the "Volcker Rule" on proprietary trading, they are not excluded from regulatory reporting on Form FR VV-1. We identify Treasury trading desks by searching for terms such as "Treasury" and "U.S. Government" in the desk name and description. As an additional check, we also review the list of exemptions from the the "Volcker Rule" on proprietary trading that the desk claims. Many Treasury trading desks claim the the "trading in domestics government obligations" and/or the "market making-related activity" exemption, which helps us narrow the sample of desks that are primarily engaged in such activities. Altogether, we identify a sample of 18 prime Treasury trading desks affiliated with 10 large banks during the sample period from 01/01/2015 to 03/31/2023. As a robustness check, we also examine a broader set of 40 desks that list Treasury securities in their desk descriptions, even if Treasury market making is not necessarily their main focus.¹

3.1.2 Internal risk limits

The "Volcker Rule" defines internal risk limits as constraints that determine the amount of risk that a trading desk is permitted to take at a point in time. Internal risk limits are not set by regulation, but rather defined by the banking entity itself for a specific trading desk as part of its internal risk management. The limits are typically expressed in terms of risk measures, such as Value-at-Risk, but may also be expressed in terms of other observable criteria, such as the number of open positions. We focus on the limits that are expressed in terms of Value-at-Risk (VaR), as all trading desks engaged in market-making related activities are required to have such limits.

To identify VaR limits, we search for limit names that include VaR. For each limit, firms report the limit size, representing an upper bound on the risk measure, as well as the actual value of usage in dollars, corresponding to the observed value of the risk measure at the end of the day. In a separate field, firms also report the VaR, defined as the risk of future financial loss in the value of the trading desk's aggregated positions at the 99% confidence level over a 1-day holding period. If several VaR limits are identified on a given desk, take the VaR limit with usage corresponding to the reported VaR measure. The *Limit usage* ratio is defined as:

$$Limit \ usage = VaR/VaR \ limit \tag{1}$$

Figure 2 shows on the left the time series of the average VaR *Limit usage* ratio. The average *Limit usage* ratio spiked in March 2020, approaching one. Despite decreasing in early April, VaR *Limit usage* at Treasury trading desks remained elevated until June.

Figure 2 plots on the right the average limit size and usage (the VaR) of the sample desks. During March 2020, the average VaR on Treasury trading desks shot up. Although firms reacted by also gradually increasing the size of the limit, the increases in limit size generally lagged the VaR increases and were not large enough fully offset the increases in VaR. This chart suggests that the

¹Many of these additional desks are repo desks, hedging desks, municipal desks, credit desks, or derivatives desks that may trade Treasury securities to support these activities.

VaR limit occasionally becomes binding, even for the average desk.

There are a couple of properties of VaR measure that are particularly relevant for our analysis. First, VaR is intrinsically counter-cyclical. Volatilities in the prices of the underlying assets translate almost one-to-one to the desk VaR, even if the desk portfolio is held constant. Second, VaR measure accounts for correlations across positions. Therefore, long and short positions of similar exposure provide off-sets in their VaR contribution. This is different from constraints that limit the gross amount of exposure. Finally, since VaR is a risk measure, it is sensitive to the amount of risks each trading position entail. For example, a Treasury bond with longer tenor contributes more to portfolio VaR than a bond with the same par value but shorter tenor. We explore such crosssectional differences in our empirical analysis.

3.1.3 Trading Positions

To meet client demand, Treasury market-making desks maintain an inventory of securities and derivatives positions. The inventory is effectively constrained by the internal risk limits at each desk. We use the size of these positions to measure the desks' market-making activities, and examine the effect of internal risk limits on Treasury market-making.

The Volcker data contains several measures of trading desk positions: the market value of long and short securities positions, and either the market or the notional value of long and short derivative positions. Prior to 12/30/2020, derivative positions were reported at the notional value, with option values delta-adjusted, and interest rate derivatives reported as 10-year bond equivalent values. Since 2020, derivative positions have been reported at market value.

We use the data to construct several inventory measures: (1) the net position, measured as the absolute value of the difference between the long and short securities and derivative positions; (2) the gross position, measured as the sum of long and short securities and derivative positions; (3) the long positions; (4) the short positions. Because these measures combine both securities and derivatives positions, we require that the derivative positions be measured as securities equivalents or at their notional value (rather than the market value). Therefore, the combined measures are only available through 12/31/2020. In addition, we construct measures based on securities positions; (1) net securities positions; (2) gross securities positions; (3) long securities positions; and (4)

short securities positions. These measures are based on the market value of securities holdings, and are available for the entire sample period.

3.2 SLR data

In addition to internal risk limits, we examine whether Treasury market-making activities are constrained by regulatory capital requirements at the bank holding company level, particularly the Supplementary Leverage Ratio (SLR). The SLR is a non-risk weighted capital requirement, measured as the ratio of a banking organization's Tier 1 capital to total leverage exposure. As such, it is particularly affected by high-volume, low-risk activities such as Treasury market intermediation.

Large U.S. banks are required to maintain an SLR of 3%; additionally, Global Systemically Important Banks (G-SIBs) are subject to an additional 2% enhanced SLR (eSLR) buffer, bringing their total SLR requirement to 5%. All but three of the sample firms, which are all foreign bank subsidiaries, are subject to the eSLR of 5%. We obtain the quarterly SLR data from the bank holding companies' consolidated financial statements (Form FR Y-9C). We measure the excess capital under the SLR rule as the *SLR distance*, the difference between the actual SLR and the minimum SLR requirement for the bank holding company in each quarter. We measure the *SLR distance* since the SLR requirement became binding on 01/01/2018.

3.3 Descriptive statistics

Table 1 shows the summary statistics for the different measures of dealer positions and internal risk limits at Treasury trading desks. The average net position including securities and derivatives is \$350.5 billion at the Treasury trading desks in the sample, and the average gross position is \$3,020.4 billion. The long and short positions are \$1,676.7 billion and \$1344.0 billion, respectively. Derivatives exposures account for most of the combined positions. Securities positions alone are notably smaller, with the average desk carrying a net securities potions of \$4.95 billion, and a gross securities position of \$26.7 billion. Variations in positions both across desks and over time are large and the distribution is skewed to the right, with the three largest desks on average accounting for almost 60% of the combined gross securities position. Therefore, we focus our analysis on the log

changes of each individual desk's positions.

The VaR usage is \$4.2 million on average, compared to the VaR limit of 12.2 million. Thus, the limit usage is about 0.3 (30%) on average. As shown in Figure 2, the limit usage varies significantly over time, approaching or exceeding one during the March 2020 Treasury market stress. There is also significant cross-sectional variation in VaR usage, which we will exploit in the analysis. For example, on March 13, just before Federal Reserve began its large scale asset purchases, the VaR limit usage on Treasury trading desks ranged between 0.15 and 1.08.

Desks' total profit and loss (PnL), that is trading revenue from existing and new trading positions, is on average \$0.98 million. Total PnL scaled by VaR is -0.34 on average.

The last row in Table 1 shows the SLR distance, which is a measure of excess capital that bank holding companies held in excess of the minimum SLR requirement. The average SLR distance is 2.4%, with a range of between 0.23% and 13%. Of note, most desks in the sample are affiliated with bank holdings companies with a minimum SLR requirement of 5%.

3.4 Fed Emergency SOMA Purchases Data

To support the smooth function of markets for Treasury securities and agency MBS, on March 15, 2020, the FOMC directed the Open Market Trading Desk (the Desk) to increase over coming months the System Open Market Account (SOMA) holdings of Treasury securities and agency mortgage-backed securities (MBS) by at least \$500 billion and at least \$200 billion, respectively. Auctions of \$40 to \$75 billion worth of Treasury securities per day were carried out from March 16th to March 23rd. On March 23rd, the FOMC further instructed the desk to increase the SOMA holdings of Treasury and agency MBS securities in the amounts needed to support the smooth functioning of both markets, essentially lending unlimited balance sheets to the market. The size of Treasury purchase operations stayed at \$75 billion per day until April 1st and only gradually decreased.

We download daily operation results from March 16, 2020 to June 11, 2020 from "Treasury Securities Operational Details" page posted by the Federal Reserve Bank of New York. Data include information on the date, CUSIP, allocated par amount and weighted average price for each securities eligible for the purchase operations. Each open market outright operation is a multipleprice, multiple securities, competitive reverse-auction (will refer it as auction for simplicity) and only primary dealers are allowed to submit offers. Duffie and Keane (2023) offers detailed descriptions of how central bank asset purchases are conducted.

There are 6,410 unique CUSIP-Date combinations with non-zero allocations during this sample period and 347 unique CUSIPs overall. For a typical CUSIP in a given day with non-zero allocation, the Desk purchases about \$133.5 million worth of the bonds from dealers.

The data does not reveal the identity of winning dealers in each auction. To infer the amount of bonds each dealer in our sample sold to SOMA, we merge this auction data with supervisory Treasury TRACE transaction reporting data. The detail of the merging process is explained in Appendix A.

We keep two versions of the merged data for our analysis. For quantity related analysis, we keep the data at dealer-day level. For price level analysis, we use transaction level (bond-dealer-day) data.

4 Dealer Inventory and VaR Limit Usage

4.1 Baseline Results

To examine the impact of dealer internal risk limits on dealer inventory, we utilize the following panel regression specification at the trading-desk level. For each Treasury market-making desk i, on day t, we estimate:

$$\Delta Log(\text{Position})_{it,t+h} = \alpha_i + \alpha_t + \beta Log(\text{Limit Usage})_{it} + c + \epsilon_{it}, \qquad (2)$$

where $\Delta Log(\text{Position})_{it,t+h} = Log(\text{Position})_{i,t+h} - Log(\text{Position})_{i,t}$, the difference in the log of desk *i*'s positions between day *t* to day t + h. Limit Usage is defined as $VaR_{it}/\overline{VaR_{it-1}}$, with $\overline{VaR_{it-1}}$ being the upper limit of VaR for desk *i*, from day t - 1. α_i and α_t are desk and time fixed-effects, respectively. We use VaR limit from day t - 1 to address the potential endogeneity of desks changing their limit size when they breach or are close to breaching it.

Our baseline results, shown in Table 2, examine the effects on net and gross positions. Since offsetting long and short positions typically have little effect on VaR calculations, the net position is likely to be most directly affected by the VaR limit. We find that Treasury market-making desks' VaR constraints have a negative and statistically significant effect on changes in net positions one day (h = 1), one week (h = 5), and two weeks ahead (h = 10). The coefficient of -0.046 in column (1) indicates that if the VaR limit usage doubles (increases by 100 percent), the change in net positions the next day decreases by 4.6 percent. The magnitude of the impact grows to 9.1 and 11.6 percent over a horizon of one week and two weeks later, respectively (columns 2 and 3). The results are similar for net securities positions, which do not include derivatives – a 100 percent increase in a desk's VaR limit usage implies a decrease in net position change the next day of 3.1 percent.

The effect of VaR limit usage on gross positions (Panel b) are also negative and statistically significant, though the magnitudes are smaller than that for net positions.

To further illustrate these results, Figure 4 shows the relationship between net (top panel) and gross (bottom panel) securities positions and Var usage. The position changes are measured over a two-week period (h = 10) and averaged across VaR usage ranks. Consistent with the regressions, the charts show that dealers are more inclined to reduce their positions as they get closer to their internal risk limit.

4.2 Inventory Aging

How do dealers go about managing their inventory subject to the VaR constraint? Do they first cut back on newer, more liquid inventory or do they first reduce older, less liquid inventory? Next, we examine whether β in equation (2) varies with inventory aging or the amount of time that securities assets (long positions) and liabilities (short positions) have been held in inventory. The analysis uses data on securities assets and liabilities, split by inventory age, which was collected until 12/31/20.

As shown in Table 3, proximity to the risk limits generally affects newer dealer inventory more strongly. As seen in panel (a), VaR limit usage has a significant effect on the one day ahead change in positions of securities assets and liabilities that is less than 30 days, but is insignificant for inventory that is more than 30 days. Over longer horizons, such as one week and two weeks ahead (panels b and c), there is statistical significance on both newer and older inventory, suggesting that constrained dealers ultimately adjust both newer and older inventory. However, the magnitudes of the impact on changes in newer inventory is generally larger. These findings suggest that dealers actively manage their inventories away from their internal risk limits by selling newer inventory first or acquiring less inventory from clients. The results for short positions (liabilities) are similar to those for long positions (assets), but smaller in magnitude.

4.3 Interaction with regulatory constraints

Given the slew of post-GFC bank regulations, a natural question is whether the effects of internal risk limits may in part reflect regulatory constraints, such as the SLR. We show this is not the case. Our findings suggest that there might be bounds in what certain policy recommendations, such as regulatory relief, could achieve in relieving dealer capacity constraints during times of market stress.

We test two regression specifications with VaR limit usage and SLR distance, the excess capital under the SLR rule:

$$\Delta Log(\text{Position})_{it,t+h} = \alpha_i + \alpha_t + \beta_1 Log(\text{Limit Usage})_{it} + \beta_2(\text{SLR dist})_{i,qtr(t)-1} + c + \epsilon_{it}, \quad (3)$$

$$\Delta Log(\text{Position})_{it,t+h} = \alpha_i + \alpha_t + \beta_1 Log(\text{Limit Usage})_{it} + \beta_2(\text{SLR dist})_{i,qtr(t)-1} + \beta_{int} Log(\text{Limit Usage}_{it}) * \text{SLR dist}_{i,qtr(t)-1} + c + \epsilon_{it},$$
(4)

In equation (3), VaR limit usage and SLR distance are both included as independent variables, while equation (4) contains the addition of an interaction term between the two. We regress on the lagged measure of SLR distance from the previous quarter to ensure there is no forward-looking information contained in the explanatory variable.

As shown in Table 4, we run regression specifications over changes in net positions (panel a) and net securities positions (panel b), as well as changes in net securities positions in 2020 (panel c) for h = 1, 5, and 10 business days ahead. In most regressions, the coefficient on VaR limit usage is negative and statistically significant. In contrast, the coefficient on lagged SLR distance is always insignificant, as well as the coefficient on the interaction term between VaR limit usage and SLR

distance. These results show that the effects of VaR limit on dealer inventory changes are not driven by the SLR.

Tables 5 and 6 provide further evidence that the effects on dealers' Treasury market-making activities are more consistent with binding VaR limits than the SLR requirement. Table 5 shows that VaR limit usage significantly affects changes in dealer gross positions, even after controlling for lagged SLR distance, which is insignificant in nearly all specifications. Table 6 documents that the effects of VaR limit usage on changes in net positions is significantly larger than its effects on changes in gross positions. In particular, if VaR limit usage were to double, dealers would be inclined to reduce the change in net positions the next day by 3.8 percent more than the change in gross positions (column 1). Over a horizon of one week and two weeks, dealers would be inclined to reduce the change in net positions by 6.8 and 8.6 percent more than that in gross positions, respectively (columns 2-3).

These results are consistent with the fact that VaR calculations usually balance long and short positions, so one would expect the net position of dealers to be more important than gross positions in keeping VaR in check. On the other hand, a size-based constraint that does not net long and short positions, such as the SLR, would have predicted the opposite, with gross positions being more important to manage than net positions.

4.4 Robustness

We show that the negative relationship between dealers' VaR limit usage and their positions in Treasury markets is robust in several dimensions.

First, the relationship holds separately for both long and short positions. As seen in Table 7, an increase in VaR limit usage has a negative and statistically significant effect on changes in long positions (panel a) and changes in short positions (panel b) for horizons of one day, one week, and two weeks ahead. The magnitude of changes in long positions is larger than that of changes in short positions, indicating that dealers reduce their long positions more than short positions as they get closer to their internal risk limit.

In addition, our findings are robust to several alternative measures of trading desks' VaR limit

usage: (a) Limit Usage (std) is a standardized measure of Limit Usage, equal to the difference between a desk's Limit Usage and its rolling mean, divided by its rolling standard deviation; (b) Limit Usage prc rank is the percentile rank of a desk's standardized Limit Usage, relative to its own distribution over the rolling window; (c) Limit Usage q2, q3, and q4 are desk-level quartile dummies, and is equal to one when a desk's Limit Usage prc rank is between 25 to 50, 50 to 75, and 75 to 100 percent, respectively. For all variables, the rolling window is the previous year.

As seen in row 1 of Table 8, the coefficient on Limit Usage (std) is negative and statistically significant for nearly all specifications, indicating that the more a desk's Limit Usage exceeds its rolling-average mean, the more the desk will reduce its positions. For instance, a one standard deviation increase in a desk's Limit Usage implies a 1.0 percent decrease in its change of net securities positions the next business day (panel b). Similarly, the coefficient on Limit Usage prc rank (row 2) is negative and statistically significant – implying that the higher the percentile rank of a desk's Limit Usage, the more it will reduce its net positions.

Limit Usage quartile dummies (rows 3-5) also have a negative relationship with positions – most of the effect is coming from when a desk's Limit Usage is in the highest quartile (q4), and the magnitudes of the coefficients on the quartile dummies generally increase with the quartile. For instance, a coefficient of -0.033 on Limit Usage q4 (panel b, column 3) indicates that a trading desk will decrease its change in net securities positions the next day by 3.3 percent when its VaR limit usage is in the highest quartile (q4) than when it's in the lowest quartile (q1). The results in Table 8 demonstrate that a desk's trading positions are not only sensitive to the level of its VaR limit usage, but also are sensitive to where its VaR limit usage falls relative to its own historical distribution.

Finally, we show that our main results hold when we expand our definition of Treasury trading desks to all desks whose name and description contains terms such as "Treasury" and "U.S. Government." The sample of Treasury trading desks increases from 18 to 40 desks affiliated with 11 large banks over our sample period 01/01/2015 to 03/31/2023. As shown in panel C of Table 7, the coefficient on log (Limit Usage) is negative and statistically significant in all specifications for the larger Treasury trading desk sample. The magnitudes of the effect is slightly smaller than for our prime market-making desk sample, though still economically meaningful – a 100 percent increase

in a desk's VaR limit usage implies a decrease in net positions change the next day of 2.5 percent (column 1).

5 Event Study: Did Fed Purchases Alleviate Dealer VaR Constraints?

To illustrate that desk level VaR limits were indeed binding in March 2020 and the months following, we conduct an event study of the effect of Federal Reserve asset purchases on dealer behavior. Specifically, we compare the likelihood, intensity and aggressiveness for dealers with different VaR constrains to sell Treasury securities to the Federal Reserve during the emergency SOMA operations from March 16, 2020 through June. The SOMA operations take bonds directly off dealers' balance sheets and allow dealers to reduce the VaR of their portfolio directly. Dealers, by having an "exit valve", should be more willing to serve their clients by taking bonds off their hands as well.

5.1 Quantities of Dealer-Sell to the Fed and VaR limit Usage

After the Federal Reserve began purchasing Treasuries on March 16, dealer inventories of Treasury securities decreased rapidly. We first document that Treasury securities declined more at dealers who were closer to their VaR constraint than at less constrained dealers. Using the Volcker data, Figure 3 contrasts two groups of desks based on the level of their VaR usage on March 13, 2020, the last business day before the announcement of the Fed the emergency SOMA operations. The more constrained desks, namely those with above median VaR usage on March 13, reduced their net positions significantly more than desks that were less constrained by their VaR limit.

Next, we examine directly whether the more constrained dealers sold more Treasury securities to the Fed SOMA portfolio. Table 9 shows results from panel regressions that study the effect of log(Limit Usage) on different measures of probability and intensity of dealer sales to the Fed the following day. The VaR usage is aggregated to the dealer level by averaging across Treasury securities trading desks. Throughout the specifications, higher usage of the VaR limit on day t - 1is associated with higher likelihood and intensity of sales by the dealer to the Fed on day t. Such effect is particularly strong in the month of March, even when both dealer and day fixed effects are accounted for.

Figure 5 illustrates this further by comparing the cumulative sales to the Fed for the VaR constrained and unconstrained dealers. The constrained dealers are taken as those with above median VaR usage on March 13, where each dealer's VaR usage is averaged across its Treasury trading desks. The chart shows that the high VaR usage dealers sold more in absolute volume to the Fed (upper panel) and as a fraction to the total amount of sales they made to customers (lower panel).

5.2 Prices of Dealer-Sell to the Fed and VaR limit Usage

Do more constrained dealers submit more aggressive offers in SOMA auctions to alleviate their VaR constraint? Recall that SOMA purchase operations are multi-price auctions where offers are ordered in prices and the best offer (lowest dealer sell price) is accepted first and other offers are accepted sequentially until an ideal total size of allocation is reached. For the same bond, different dealers could be offering at different prices. We hypothesize that dealers closer to their VaR limits would be willing to offer more aggressively and accept lower prices in auctions. Table 10 shows results from regressions that link dealer offer aggressiveness to their VaR limit usage in the day prior. Dealer offer aggressiveness is measured as the difference between average price accepted in an auction and the dealer's actual sell price to the Fed ($\Delta p = \text{Avg}$ Price in Auction – Price_{dealer sell}). ² We show that dealers that are closer to their VaR limit offer more aggressively in auctions, and the effect is stronger in March 2020.

Moreover, since the price risk of Treasury bonds scales with duration, bonds with more years to mature (remaining life) contribute more to the portfolio VaR. VaR-constrained dealers are therefore incentivized to get rid of long maturity bonds first. The granularity of the data in Table 10 allow us to test this hypothesis directly. Indeed, for bonds with more years to mature, dealer' offer aggressiveness is more sensitive to its internal risk limit usage. This differentiation across bonds implies that bonds contribute differently to dealers' balance sheets constraint, and the constraint

²While dealer offers are not disclosed to the public, we infer dealer offers through reported transaction prices. These transactions are matched to the SOMA auctions using method documented in Appendix A.

at work is therefore not "risk blind".

5.3 Dealer Liquidity Provision to Clients and Fed Purchases

Does the ability to off-load bonds to the Federal Reserve translate to more dealer intermediation in the form of liquidity provision to clients? Since client order flow was overwhelmingly in the direction of sells, dealer liquidity supply during this period would be in the form of dealer buys from clients. In Table 11 we regress dealer buy volume on its prior day VaR limit usage and dealer sales to SOMA on the same day. The panel regression results suggest that higher VaR limit usage the day prior is associated with less client intermediation— more constrained dealers buy less from clients. However, controlling for VaR limit usage, higher volume of dealer sales to SOMA is associated with higher volume of buys from clients. Having the Fed SOMA portfolio as an exit valve allows primary dealers to supply more liquidity to their clients. Not only are dealers able to buy more bonds from clients for bonds that are in the SOMA auction lists, dealers are able to buy more bonds not on the SOMA operation list from their clients, as their balance sheets become less constrained due to SOMA operations. Moreover, the effect of SOMA sales on dealer liquidity provision is primarily coming from the month of March, when market stress was the highest.

Our event study results suggest that the SOMA emergency purchase program was effective in providing liquidity to the market since the Fed was lending its balance sheets to primary dealers, who became less constrained themselves by selling their Treasury holdings directly into the SOMA portfolio and making room in their portfolio to provide more immediacy to clients.

6 Market Liquidity and VaR Usage

6.1 Dealer Profitability and VaR Usage

How does dealer's risk aversion affect market liquidity? In Section 4, we documented that dealers adjust the size of their inventory as they approach the VaR limit. In this section, we further show that dealers adjust the pricing of their liquidity provision, requiring higher compensation if their VaR constraint is more binding. In particular, we estimate:

$$\frac{1}{h} \cdot \sum_{k=1}^{h} \left[\frac{\mathrm{P\&L}}{\mathrm{VaR}} \right]_{i,t+k} - \left[\frac{\mathrm{P\&L}}{\mathrm{VaR}} \right]_{i,t} = \alpha_i + \alpha_t + \beta Log(\mathrm{Limit} \; \mathrm{Usage})_{it} + c + \epsilon_{it},$$

where the dependent variable is the average change in a desk's profitability (scaled by VaR) over the next h business days. Higher future profitability per unit of risk can be achieved by dealers either setting higher prices for providing immediacy (by taking bonds into their inventory) or being more selective with new trades and only engage capital when the trade is sufficiently profitable.

The results in Table 12 show that as dealers' VaR limit usage increases, they seek greater compensation, in terms of trading profitability, for taking on the increased risk of breaching their internal risk limits. In other words, a desk's average future P&L scaled by its VaR increases with the tightness of its VaR constraint. The effects persist up to two weeks later (Panel c) and are statistically significant. Notably, the effect is driven entirely by P&L generated from desks' *new* trading positions (column 2), rather than their existing positions (column 3). A 100 percent increase in a desk's VaR limit usage corresponds with dealers requiring 7.4 percentage points higher P&L per unit of VaR to take on new positions over the next business day (column 1). P&L from new trades are mostly driven by markups, commissions clients paid, whereas P&L from existing trades could be driven by existing positions' exposures to different risk factors. Therefore, our results on dealer profitability is indicative of dealer pricing behavior, rather than their ability to predict future market movement. Since the flip side of dealer profitability is client transaction cost, this result offers a direct link between dealer capacity constraints and client transaction costs.

6.2 Aggregate Market Liquidity and VaR Usage Index

Taking a macro view, we examine next the implications of trading desks' VaR limit usage for aggregate Treasury market liquidity. We first construct an aggregate measure, or an index, of dealers' Treasury market VaR constraints from desk-level VaR limit usage, weighted by a desk's market share of customer transaction volumes over the past 20 business days (w_{it}) :

Limit Usage Index_t =
$$\sum_{i} \mathbf{w}_{it} \cdot Log(\text{Limit Usage})_{it}$$
, (6)

where

$$\mathbf{w}_{it} = \frac{\sum_{k=1}^{20} \text{Cust Trans Volumn}_{i,t-k}}{\sum_{i} \left[\sum_{k=1}^{20} \text{Cust Trans Volumn}_{i,t-k}\right]}.$$
(7)

It is well documented that Treasury market liquidity is highly sensitive to interest rate volatility. Our limit usage index is also positively correlated with interest rate volatility. To isolate the independent effect of limit usage on market liquidity, we first regress our limit usage index on a market-implied measure of bond market volatility, the MOVE index, and extract the residual. We then conduct quantile regressions of Treasury market illiquidity at the τ th percentile on the Limit Usage Index residual, while controlling for interest rate volatility (MOVE index):

$$Q(\tau)_{\text{Illiquidity},t} = \alpha + \beta_{\tau} (\text{Limit Usage Index residual})_t + \gamma (\text{MOVE index})_t + \epsilon_t, \tag{8}$$

Table 13 shows quantile regression results for two Treasury market illiquidity measures. In Panel (a), the measure is the average absolute residuals that result from fitting a smooth Treasury yield curve to the cross-section of yields every day.³ This measure of market illiquidity is used in D'Amico and King (2013) and Hu, Pan, and Wang (2013). The impact of VaR constraint on Treasury market illiquidity is skewed and more pronounced for higher percentiles of market illiquidity, consistent with Duffie, Fleming, Keane, Nelson, Shachar, and Van Tassel (2023), who show that market illiquidity changes in character at tail levels of illiquidity, becoming more sensitive to measures of dealer intermediation capacity. For instance, column 4 shows that for the 99th percentile, a one standard deviation increase in the Limit Usage Index residual corresponds with a 0.24 standard deviation increase in Treasury market illiquidity. Pseudo R^2 also increases with market illiquidity percentile, nearly doubling from the 50th percentile to the 99th percentile regression.

Interestingly, unlike Duffie, Fleming, Keane, Nelson, Shachar, and Van Tassel (2023), who only find significant relationship between their measure of dealer capacity and Treasury market illiquidity at extreme tails, we find significant relationship between dealer VaR constraints and Treasury market illiquidity both on average, through an OLS regression (column 1) and in the median, through quantile regression at 50% (column 2).

 $^{^{3}}$ We are grateful to Federal Reserve Board staff for providing the data based on internal calculations. Sample used in the yield curve estimation includes Treasury securities maturing in between 2 and 10 years, excluding on-the-run and first off-the-run securities.

In Panel (b), the measure of Treasury market illiquidity is the on-the-run yield premia, defined as the spread between the 10-year on-the-run and off-the-run Treasury yield. Similarly, we find that the relationship between VaR constraints and Treasury market illiquidity is statistically significant for both average and tail measures of illiquidity. Both the coefficient on the Limit Usage Index residual and the pseudo R^2 increases with illiquidity percentile. Overall, our findings support the hypothesis that the tightness of internal risk limits at dealers adversely affect market liquidity, and particularly so in times of market stress.

7 Conclusion

We provide direct, trading desk-level evidence that dealer internal risk limits in the form of VaR limits limit dealer willingness to take risk in the Treasury market, especially in times of market stress. Dealers avoid taking on additional inventory and require higher compensation per unit of risk when their trading positions are closer to their VaR limits. In March 2020, many primary dealer Treasury securities trading desks either breached or were close to breaching their VaR limit, severely limiting these desks' ability to provide liquidity to clients. We show that dealers that were closer to their VaR limit were more likely to sell Treasury securities to the Federal Reserve during the emergency SOMA operations authorized in late March 2020. All else equal, dealers that sold more to the Federal Reserve were able to provide more liquidity to clients, both in bonds that were on the list of Federal Reserve purchases and bonds outside the lists, suggesting that the bond purchase program was effective in alleviating dealer VaR constraint by taking bonds directly off their balance sheets.

Our work has important policy implications. We should be careful attributing "risk averse" dealer behavior to financial regulations such as the SLR. Self-imposed internal risk limits could be lurking in the background. Such insights could matter in assessing various options of structural reforms that aim at improving Treasury market liquidity. Deregulation could be limited in its potency in removing dealer balance sheet rigidity. Treasury market reforms that focus on reducing market's reliance on dealers' balance sheets (such as encouraging all-to-all trading) could potentially be more effective. Our research also offers insights into what market intervention methods are

most useful in addressing market functioning problems during market stress. We show that in a crisis where intermediaries are risk averse, policies that address the funding costs of intermediaries wouldn't be as effective as policies that remove risks from intermediary balance sheets directly. More work needs to be done to understand how different constraints that dealers face affect their behavior and how these constraints interact.

References

- Adrian, Tobias, and Hyun Song Shin, 2013, Procyclical Leverage and Value-at-Risk, The Review of Financial Studies 27, 373–403.
- Amihud, Yakov, and Haim Mendelson, 1980, Dealership market: Market-making with inventory, Journal of Financial Economics 8, 31–53.
- Anderson, Christopher S., David C. McArthur, and Ke Wang, 2023, Internal risk limits of dealers and corporate bond market making, *Journal of Banking & Finance* 147, 106653.
- Aragon, George O., and Philip E. Strahan, 2012, Hedge funds as liquidity providers: Evidence from the Lehman bankruptcy, *Journal of Financial Economics* 103, 570–587.
- Boyson, Nicole M., Christof W. Stahel, and Rene M. Stulz, 2010, Hedge fund contagion and liquidity shocks, 65, 1789–1816.
- Breckenfelder, J, and V Ivashina, 2021, Bank balance sheet constraints and bond liquidity, Discussion paper.
- Duffie, Darrell, 2018, Financial regulatory reform after the crisis: An assessment, *Management Science* 64, 4835–4857.
- ———, 2023, Resilience redux in the us treasury market, Discussion paper.
- , Michael J. Fleming, Frank M. Keane, Claire Nelson, Or Shachar, and Peter Van Tassel, 2023, Dealer capacity and u.s. treasury market functionality, .
- Duffie, Darrell, and Frank M. Keane, 2023, Market-function asset purchases, Staff Reports.
- D'Amico, Stefania, and Thomas B King, 2013, Flow and stock effects of large-scale treasury purchases: Evidence on the importance of local supply, *Journal of Financial Economics* 108, 425–448.
- Falato, Antonio, Itay Goldstein, and Ali Hortaçsu, 2021, Financial fragility in the covid-19 crisis:The case of investment funds in corporate bond markets, *Journal of Monetary Economics*.

- Goldberg, Jonathan, 2020, Dealer inventory constraints during the covid-19 pandemic: Evidence from the treasury market and broader implications, Discussion paper.
- He, Zhiguo, Stefan Nagel, and Zhaogang Song, 2021, Treasury inconvenience yields during the covid-19 crisis, *Journal of Financial Economics*.
- Hu, Grace Xing, Jun Pan, and Jiang Wang, 2013, Noise as information for illiquidity, *The Journal of Finance* 68, 2341–2382.
- Kruttli, Mathias, Phillip Monin, Lubomir Petrasek, and Sumudu W. Watugala, 2021, LTCM redux?: Hedge fund Treasury trading and funding fragility, Discussion paper.
- Shleifer, Andrei, and Robert W. Vishny, 1997, The limits of arbitrage, *The Journal of Finance* 52, 35–55.
- Vissing-Jorgensen, Annette, 2021, The Treasury market in spring 2020 and the response of the Federal Reserve, *Journal of Monetary Economics* 124, 19–47.

Appendix

A Matching SOMA Operation data and Treasury TRACE

We match Treasury transactions in the supervisory TRACE to the SOMA auctions data in March through June, 2020 to identify the quantities and prices of Treasuries that primary dealers sold to the Federal Reserve Board.

We take transactions that are "Dealer sell to customer" in principal capacity by a primary dealer and a par amount no less than 1/2 of a million to SOMA auction operations on the same day, for the same CUSIP. The match is done sequentially from level 1 through level 3, with level 1 being the most precise and level 3 being the least.

- 1. In level 1 match, we require the size of the transaction to match exactly the size of the bond awarded in the SOMA auction. We further require that the timestamp of the trade to be within [-10, 30] minutes of the auction timestamp. When there are multiple trades that match the auction record, we pick the trade that has a timestamp closer to the auction time. Roughly a quarter of all auctions can be matched uniquely to a transaction in TRACE.
- 2. In Level 2 match, we allow TRACE transaction quantities to be smaller than the auction sizes and several transactions add up to the sizes of the auction. We find that roughly 37% of all SOMA auctions can be matched to a series of TRACE transactions, whose par amount add up exactly to the quantity awarded in the auction, and another 30% can be matched to a series of transactions that add up to a quantity smaller than the auction amount.
- 3. In Level 3 match, we take the set of transactions of the same CUSIP and sort them by transaction prices, we add up the trade volume from the one with the highest prices down until we exhaust the auction volume. Only 7% of SOMA auctions are matched this way.

The end results of the match are trades that are dealer sell to the Fed during the period from March 16 to June 11th, 2020. We then match this data to our Volcker data at dealer-day level for analysis.

B Tables

	mean	sd	p25	p50	p75	count
Net position	350.5	1209.1	0.34	3.24	40.5	31641
Gross position	3020.4	5475.9	10.0	340.7	4097.2	31641
Long position	1676.7	2987.1	5.24	182.4	2145.5	31657
Short position	1344.0	2609.4	4.72	157.1	1740.2	31686
Net securities position	4.95	9.01	0.23	2.15	6.41	43310
Gross securities position	26.7	39.2	2.25	13.8	30.0	43310
Long securities position	15.0	21.4	1.15	7.98	17.8	43313
Short securities position	11.7	18.9	0.67	4.73	13.2	43323
VaR	4.20	6.25	0.82	2.26	4.73	42483
VaR limit	12.2	12.1	5	10	15	41904
Limit Usage	0.30	0.20	0.16	0.26	0.40	41814
Total PnL	0.98	4.27	-0.068	0.21	1.21	45840
Total PnL / VaR	-0.34	37.1	-0.047	0.18	0.55	42378
SLR dist	2.40	2.37	1.30	1.61	2.36	29863

Table 1: Summary Statistics

Table 1 shows summary statistics for the positions variables (in billions of dollars), VaR and VaR limit (in millions of dollars), Limit Usage (defined as the ratio of VaR relative to VaR limit from the day before), total profit and loss (PnL) in millions of dollars, total PnL scaled by VaR, and SLR distance (in percent). Gross (securities) positions are the sum of long and short (securities) positions. Net (securities) positions are the absolute value of the difference between long and short (securities) positions. Positions include cash and derivative assets. Securities positions include only cash assets. The sample period is 1/1/15-12/31/20 for positions and 1/1/15-3/31/23 for securities positions.

Table 2: Dealer Desk Positions and VaR Constraints (Baseline)

			r			
	N	et positions	5	Net se	curities posi	tions
	(1)	(2)	(3)	(4)	(5)	(6)
	h=1	h=5	h=10	h=1	h=5	h=10
log(Limit Usage)	-0.046^{***} (0.017)	-0.091*** (0.021)	$\begin{array}{c} -0.116^{***} \\ (0.023) \end{array}$	$\begin{array}{c} -0.031^{***} \\ (0.011) \end{array}$	-0.058^{***} (0.014)	-0.070^{***} (0.016)
Desk FE	Yes	Yes	Yes	Yes	Yes	Yes
Time FE	Yes	Yes	Yes	Yes	Yes	Yes
R^2	0.131	0.127	0.144	0.129	0.125	0.132
N	12045	11986	11915	16838	16786	16721
		(b) Gro	ss positions			
	(Gross positi	ons	Gross	securities p	ositions
	(1)	(2)	(3)	(4)	(5)	(6)
	h=1	h=5	h=10	h=1	h=5	h=10
ln_limit_usage_ratio2	-0.006^{***}	-0.017^{***}	-0.024^{***}	-0.007^{***}	-0.020***	-0.030***
	(0.001)	(0.002)	(0.003)	(0.002)	(0.003)	(0.003)
Desk FE	Yes	Yes	Yes	Yes	Yes	Yes
Time FE	Yes	Yes	Yes	Yes	Yes	Yes
R^2	0.136	0.146	0.154	0.132	0.153	0.154
N	12045	11986	11915	16838	16786	16721

(a) Net positions

Table 2 shows the effect of log(Limit Usage) on the log change in net positions (panel a) and on the log change in gross positions (panel b). Net and gross positions (columns 1 through 3) include cash and derivative assets. Net and gross securities positions (columns 4 through 6) include only cash assets. Net (securities) positions are the absolute value of the difference between long and short (securities) positions. Gross (securities) positions are the sum of long and short (securities) positions. Limit Usage is defined as the ratio of VaR relative to VaR limit from the day before. Regression results are shown for positions change over h = 1, 5, and 10 business days. The sample period is $1/1/15 \cdot 12/31/20$ for positions (columns 1 through 3) and $1/1/15 \cdot 3/31/23$ for securities positions (columns 4 through 6). All regressions include desk and time fixed effects. Standard errors are robust.

Table 3: Effect of VaR Constraints on Dealer Long Positions (Assets) and Short Positions (Liabilities) of Different Vintages

	(a) h=1						
	$(1) \\ Assets \\ (\leq 30 \text{ day})$	(2) Assets (>30 day)	$\begin{array}{c} (3) \\ \text{Liab} \\ (\leq 30 \text{ day}) \end{array}$	(4) Liab (>30 day)				
log(Limit Usage)	-0.010^{***} (0.003)	-0.002 (0.002)	-0.008** (0.003)	-0.003 (0.003)				
R^2 N	$0.152 \\ 11428$	$0.242 \\ 11426$	$0.146 \\ 11434$	0.208 11323				
(b) h=5								
	$(1) \\ Assets \\ (\leq 30 \text{ day})$	(2) Assets (>30 day)	(3)Liab ($\leq 30 \text{ day}$)	(4) Liab (>30 day)				
log(Limit Usage)	-0.034^{***} (0.005)	-0.018^{***} (0.005)	-0.013** (0.006)	-0.017** (0.007)				
R^2 N	$0.168 \\ 11376$	$0.150 \\ 11374$	$0.150 \\ 11382$	$0.165 \\ 11260$				
(c) $h=10$								
	$(1) \\ Assets \\ (\leq 30 \text{ day})$	(2) Assets (>30 day)	(3)Liab ($\leq 30 \text{ day}$)	(4) Liab (>30 day)				
log(Limit Usage)	-0.047^{***} (0.006)	-0.029^{***} (0.008)	-0.029^{***} (0.007)	-0.020** (0.010)				
R^2 N	$0.176 \\ 11309$	$0.151 \\ 11309$	$0.146 \\ 11315$	$0.160 \\ 11183$				

Table 3 shows the effect of log(Limit Usage) on the log change in securities assets (columns 1 to 2) and securities liabilities (columns 3 to 4), separated by inventory age. Limit Usage is defined as the ratio of VaR relative to VaR limit from the day before. Panels (a), (b) and (c) show regression results for securities change over h = 1, 5, and 10 business days, respectively. The sample period is 1/1/15-12/31/20. All regressions include desk and time fixed effects. Standard errors are robust.

Table 4: Regulatory Constraints and Net Positions

	(a) I	Net positic	ons			
	h=	=1	h=	5	h=	10
	(1)	(2)	(3)	(4)	(5)	(6)
log(Limit Usage)	-0.029 (0.027)	-0.045 (0.040)	-0.088*** (0.033)	-0.111^{**} (0.050)	-0.129^{***} (0.035)	-0.175^{***} (0.052)
lagged SLR dist	-0.002 (0.025)	0.012 (0.033)	-0.007 (0.032)	0.013 (0.042)	-0.020 (0.036)	0.021 (0.046)
log(Limit Usage) \times lagged SLR dist	× ,	0.007 (0.012)		0.010 (0.016)		0.021 (0.016)
R^2	0.121	0.121	0.127	0.127	0.140	0.140
N	6160	6160	6113	6113	6057	6057
	(b) Net se	ecurities p	ositions			
	h=	=1	h=	=5	h=	=10
	(1)	(2)	(3)	(4)	(5)	(6)
log(Limit Usage)	-0.037^{**} (0.015)	-0.033 (0.022)	-0.057^{***} (0.019)	-0.070^{***} (0.027)	-0.072^{***} (0.022)	-0.071^{*} (0.031)
lagged SLR dist	0.001	-0.000	-0.010	-0.005	-0.019	-0.019
$\log(\text{Limit Usage}) \times \text{lagged SLR dist}$	(0.013)	(0.014) -0.002	(0.016)	$(0.018) \\ 0.006$	(0.017)	(0.019) -0.000
$\log(\text{Limit Usage}) \times \text{lagged SLR dist}$		(0.002)		(0.000)		(0.010)
R^2	0.122	0.122	0.116	0.116	0.115	0.115
Ν	10947	10947	10907	10907	10857	10857
(c) Net secu	rities posi	tions 2020			
	ł	n=1	h	=5	h=	10
	(1)	(2)	(3)	(4)	(5)	(6)
log(Limit Usage)	-0.061	-0.063	-0.136***	-0.139**	-0.135**	-0.120
	(0.039)	(0.052)	(0.051)	(0.070)	(0.055)	(0.076)
lagged SLR dist	-0.003	-0.003	-0.128	-0.128	-0.114	-0.117
$\log(\text{Limit Usage}) \times \text{lagged SLR dist}$	(0.092)	(0.091) 0.001	(0.122)	(0.122) 0.002	(0.135)	(0.135) -0.008
$\log(\text{DHILL Osage}) \times \log(\text{DLL Osage})$	J	(0.001)		(0.002)		(0.024)
R^2	0.124	0.124	0.113	0.113	0.118	0.118
N	2232	2232	2228	2228	2223	2223

Table 4 shows the effects of dealer regulatory constraint, SLR distance, and log(Limit Usage) on the log change in net positions (panel a), log change in net securities positions (panel b), and the log change in net securities positions in 2020 (panel c). Net (securities) positions are the absolute value of the difference between long and short (securities) positions. Net positions include cash and derivative assets. Net securities positions include only cash assets. Lagged SLR distance is defined to be the previous quarter's difference between a firm's SLR and the minimum SLR requirement as specified in the Basel III accord. Data for the SLR starts in 1/1/18. Limit Usage is defined as the ratio of VaR relative to VaR limit from the day before. Regression results are shown for positions change over h = 1, 5, and 10 business days. The sample period for regressions in panel (a) is 1/1/18-12/31/20; in panel (b) is 1/1/18-3/31/23; and in panel (c) is 1/1/20-12/31/20. All regressions include desk and time fixed effects. Standard errors are robust.

Table 5: Regulatory Constraints and Gross Positions

	h=	-1	h=	=5	h=	=10
	(1)	(2)	(3)	(4)	(5)	(6)
log(Limit Usage)	-0.010***	-0.009**	-0.029***	-0.031***	-0.043***	-0.045***
	(0.003)	(0.004)	(0.005)	(0.006)	(0.006)	(0.008)
lagged SLR dist	0.000 (0.002)	-0.002 (0.002)	0.001 (0.003)	0.002 (0.004)	0.002 (0.004)	0.005 (0.006)
$\log(\text{Limit Usage}) \times \text{lagged SLR dist}$	(0.002)	(0.002)	(0.005)	(0.004) 0.001	(0.004)	0.001
		(0.001)		(0.002)		(0.002)
R^2	0.122	0.122	0.123	0.123	0.133	0.133
Ν	6160	6160	6113	6113	6057	6057
	(b) Gross	securities p	ositions			
	h=	:1	h=	=5	h=	=10
	(1)	(2)	(3)	(4)	(5)	(6)
log(Limit Usage)	-0.009***	-0.011**	-0.029***	-0.033***	-0.040***	-0.043***
	(0.003)	(0.005)	(0.004)	(0.007)	(0.005)	(0.008)
lagged SLR dist	-0.001	-0.001	-0.002	-0.000	-0.010	-0.009
	(0.004)	(0.005)	(0.006)	(0.006)	(0.007)	(0.008)
$\log(\text{Limit Usage}) \times \text{lagged SLR dist}$		0.001 (0.002)		0.002 (0.003)		0.002 (0.003)
R^2	0.127	0.127	0.153	0.153	0.152	0.152
N N	10.127 10947	10.127 10947	10.155 10907	10.135 10907	10857	10.152 10857
()	c) Gross sec	urities posi	itions 2020			
	h=1		h=	=5	h=	10
	(1)	(2)	(3)	(4)	(5)	(6)
log(Limit Usage)	-0.020***	-0.014*	-0.047***	-0.035**	-0.057***	-0.046***
	(0.007)	(0.008)	(0.011)	(0.014)	(0.013)	(0.017)
lagged SLR dist	-0.015	-0.016	-0.034	-0.036	-0.069**	-0.071**
log(Limit Hoome) v logonal CLD l'	(0.016)	(0.016)	(0.027)	(0.027)	(0.031)	(0.031) -0.005
log(Limit Usage) \times lagged SLR dist		-0.003 (0.002)		-0.006 (0.004)		(0.005)
R^2	0.133	0.134	0.172	0.173	0.181	0.181
N	2232	2232	2228	2228	2223	2223

Table 5 shows the effects of dealer regulatory constraint, SLR distance, and log(Limit Usage) on the log change in gross positions (panel a), log change in gross securities positions (panel b), and the log change in gross securities positions in 2020 (panel c). Gross (securities) positions are the sum of long and short (securities) positions. Gross positions include cash and derivative assets. Gross securities positions include only cash assets. Lagged SLR distance is defined to be the previous quarter's difference between a firm's SLR and the minimum SLR requirement as specified in the Basel III accord. Data for the SLR starts in 1/1/18. Limit Usage is defined as the ratio of VaR relative to VaR limit from the day before. Regression results are shown for positions change over h = 1, 5, and 10 business days. The sample period for regressions in panel (a) is 1/1/18-12/31/20; in panel (b) is 1/1/18-3/31/23; and in panel (c) is 1/1/20-12/31/20. All regressions include desk and time fixed effects. Standard errors are robust.

	Net	/Gross posi	tions	Net/Gross securities positions			
	(1) $h=1$	(2) h=5	(3) h=10	(4) h=1	(5) h=5	(6) h=10	
log(Limit Usage)	-0.038^{**} (0.017)	-0.068^{***} (0.021)	-0.086^{***} (0.022)	-0.023^{**} (0.011)	-0.035^{**} (0.014)	-0.038^{**} (0.015)	
$\frac{R^2}{N}$	$0.132 \\ 12045$	$0.128 \\ 11986$	$0.145 \\ 11915$	$0.130 \\ 16838$	$0.125 \\ 16786$	$0.132 \\ 16721$	

Table 6: Net positions relative to gross positions

Table 6 shows the effect of log(Limit Usage) on the log change in the ratio of net positions to gross positions (columns 1 through 3) and on the log change in the ratio of net securities positions to gross securities positions (columns 4 through 6). Net (securities) positions are the absolute value of the difference between long and short (securities) positions. Gross (securities) positions are the sum of long and short (securities) positions. Net (gross) positions include cash and derivative assets. Net (gross) securities positions include only cash assets. Limit Usage is defined as the ratio of VaR relative to VaR limit from the day before. Regression results are shown for positions change over h = 1, 5, and 10 business days. The sample period is 1/1/15-12/31/20 for positions (columns 1 through 3) and 1/1/15-3/31/23 for securities positions (columns 4 through 6). All regressions include desk and time fixed effects. Standard errors are robust.

Table 7: Robustness

	L	ong position	S	Long	securities po	ositions			
	(1) h=1	$\begin{array}{c} (2) \\ h=5 \end{array}$	(3) h=10	(4) h=1	$(5) \\ h=5$	(6) h=10			
1 (T · · · · T)		-							
log(Limit Usage)	-0.007^{***} (0.002)	-0.020^{***} (0.003)	-0.030^{***} (0.004)	-0.008^{***} (0.002)	-0.025^{***} (0.003)	-0.034^{***} (0.004)			
R^2	0.133	0.143	0.154	0.134	0.153	0.161			
N	12045	11986	11915	16838	16786	16721			
(b) Short Positions									
	S	Short positio	ons	Short	securities p	ositions			
	(1)	(2)	(3)	(4)	(5)	(6)			
	h=1	h=5	h=10	h=1	h=5	h=10			
log(Limit Usage)	-0.005***	-0.009***	-0.015***	-0.005*	-0.013***	-0.023***			
	(0.001)	(0.003)	(0.003)	(0.003)	(0.004)	(0.005)			
R^2	0.135	0.142	0.144	0.131	0.140	0.136			
N	12047	11989	11921	16825	16773	16706			
(c) Net Positions - all UST market-making desks									
	ľ	Net positions	3	Net s	ecurities pos	itions			
	(1) h=1	(2) h=5	(3) h=10	(4) h=1	(5) h=5	(6) h=10			

(a) Long Positions

		Net position	ıs	Net s	Net securities positions			
	(1)	(2)	(3)	(4)	(5)	(6)		
	h=1	h=5	h=10	h=1	h=5	h=10		
log(Limit Usage)	-0.025^{**}	-0.053^{***}	-0.075^{***}	-0.020***	-0.045^{***}	-0.063^{***}		
	(0.011)	(0.014)	(0.016)	(0.008)	(0.010)	(0.012)		
$\frac{R^2}{N}$	$0.066 \\ 23471$	$0.064 \\ 23349$	$0.067 \\ 23201$	$0.062 \\ 32560$	$0.063 \\ 32438$	$0.069 \\ 32292$		

Table 7 shows the effect of log(Limit Usage) on the log change in long positions (Panel a) and on the log change in short positions (Panel b). Panel c shows the effect of log(Limit Usage) on the log change in net positions (columns 1 through 3) and on the log change in net securities positions (columns 4 through 6) for all UST market-making desks. Net (securities) positions are the absolute value of the difference between long and short (securities) positions. Positions include cash and derivative assets. Securities positions include only cash assets. Limit Usage is defined as the ratio of VaR relative to VaR limit from the day before. Regression results are shown for positions change over h = 1, 5, and 10 business days. The sample period is 1/1/15-12/31/20 for positions (columns 1 through 3) and 1/1/15-3/31/23 for securities positions (columns 4 through 6). All regressions include desk and time fixed effects. Standard errors are robust.

	h=1				h=5			h=10		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	
Limit Usage (std)	-0.007 (0.007)			-0.017^{*} (0.009)			-0.038^{***} (0.010)			
Limit Usage prc rank	· · · ·	-0.046 (0.030)		· · · ·	-0.099^{***} (0.037)			-0.153^{***} (0.039)		
Limit Usage q2		. ,	-0.028 (0.023)		. ,	-0.073^{**} (0.030)			-0.078^{**} (0.032)	
Limit Usage q3			-0.019 (0.024)			-0.085 ^{***} (0.030)			-0.079^{**} (0.032)	
Limit Usage q4			-0.039 (0.024)			-0.062** (0.030)			-0.108*** (0.032)	
R^2	0.131	0.133	0.131	0.125	0.127	0.125	0.143	0.145	0.143	
Ν	11853	11673	11882	11794	11614	11823	11723	11543	11752	

Table 8: Robustness - Limit Usage Rank Transformation

(a) Net positions

(b) Net securities positions

	h=1			h=5			h=10	
(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
-0.010** (0.005)			-0.016^{**} (0.006)			-0.023^{***} (0.007)		
()	-0.044^{**} (0.018)		~ /	-0.073^{***} (0.024)			-0.093^{***} (0.026)	
		-0.007 (0.015)		· · ·	-0.041^{**} (0.020)		· · ·	-0.064^{***} (0.022)
		-0.011			-0.028			-0.032 (0.022)
		-0.033^{**} (0.015)			-0.059^{***} (0.019)			-0.078^{***} (0.021)
0.129	0.131	0.128	0.124	0.125	0.124	0.131	0.131	$0.131 \\ 16713$
	-0.010** (0.005)	(1) (2) -0.010** (0.005) -0.044** (0.018) 0.129 0.131	$\begin{array}{ c c c c c c c c c c c c c c c c c c c$	$\begin{array}{ c c c c c c c c c c c c c c c c c c c$	$\begin{array}{ c c c c c c c c c c c c c c c c c c c$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$

Table 8 shows the effect of alternative measures of Limit Usage on the log change in net positions (Panel a) and the log change in net securities positions (Panel b). Net (securities) positions are the absolute value of the difference between long and short (securities) positions. Net positions include cash and derivative assets. Net securities positions include only cash assets. Limit Usage (std) is a standardized measure of Limit Usage, equal to the difference between a desk's Limit Usage and its rolling mean, divided by its rolling standard deviation. Limit Usage prc rank is the percentile rank of a desk's standardized Limit Usage, relative to its own distribution over the rolling window. Limit Usage q2, q3, and q4 are desk-level quartile dummies, and is equal to one when a desk's Limit Usage prc rank is between 25 to 50, 50 to 75, and 75 to 100 percent, respectively. For all variables, the rolling window is the previous year. Regression results are shown for positions change over h = 1, 5, and 10 business days. The sample period is 1/1/15-12/31/20 for positions (panel a) and 1/1/15-3/31/23 for securities positions (panel b). All regressions include desk and time fixed effects. Standard errors are robust.

	Table 9:	Dealer	Sales to	SOMA
--	----------	--------	----------	------

			Р	anel A	unel A			
	Probability of Sale				Log Amount	of Sale		
	(1)	(2)	(3)	(4)	(5)	(6)		
$\log(\text{Limit Usage})_{t-1}$	0.164^{***} (0.022)	0.121^{***} (0.042)		0.331^{***} (0.113)	0.693^{***} (0.243)			
march= $0 \times \log(\text{Limit Usage})_{t-1}$	()	()	0.025 (0.036)		· · /	0.020 (0.188)		
march= $1 \times \log(\text{Limit Usage})_{t-1}$			(0.000) (0.127^{***}) (0.046)			(0.100) (0.330) (0.214)		
cons	0.936^{***} (0.013)	0.903^{***} (0.032)	(0.010) 0.845^{***} (0.027)	6.269^{***} (0.075)	6.517^{***} (0.180)	6.097^{***} (0.126)		
Dealer FE Time FE	Yes	Yes	Yes Yes	Yes	Yes	Yes Yes		
R ² N	$0.296 \\ 625$	$\begin{array}{c} 0.414 \\ 625 \end{array}$	$\begin{array}{c} 0.637 \\ 625 \end{array}$	$\begin{array}{c} 0.515\\518\end{array}$	$0.351 \\ 518$	$\begin{array}{c} 0.760 \\ 518 \end{array}$		
	Panel B Sales to SOMA/Total Sales to Cust. Sales to SOMA/Total Buys from							
	(1)	(2)	(3)	(4)	(5)	(6)		
$\log(\text{Limit Usage})_{t-1}$	0.059^{***} (0.010)	0.098^{***} (0.023)		0.104^{***} (0.016)	0.137^{***} (0.036)			
March= $0 \times \log(\text{Limit Usage})_{t-1}$	(0.010)	(0.020)	0.024 (0.022)	(0.010)	(0.000)	0.037 (0.036)		
March= $1 \times \log(\text{Limit Usage})_{t-1}$			(0.022) 0.056^{**} (0.026)			(0.080) 0.082^{**} (0.041)		
cons	$\begin{array}{c} 0.234^{***} \\ (0.010) \end{array}$	$\begin{array}{c} 0.264^{***} \\ (0.019) \end{array}$	$\begin{array}{c} (0.020) \\ 0.213^{***} \\ (0.018) \end{array}$	$\begin{array}{c} 0.330^{***} \\ (0.019) \end{array}$	0.355^{***} (0.031)	$\begin{array}{c} (0.011) \\ 0.285^{***} \\ (0.030) \end{array}$		
Dealer FE Time FE	Yes	Yes	Yes Yes	Yes	Yes	Yes Yes		
				1				

Table shows the effect of log(Limit Usage) on the probability and quantity of dealer sale to the Fed SOMA portfolio between March 16 and June 11th, 2020. In Panel A, the dependant variables for Columns 1 through 3 are the probability of sales and the dependant variables for Column 4 through 6 are the log volume of sales. In Panel B, the dependant variable for Columns 1 through 3 are the ratio of dealer sales to SOMA to total dealer sales to customer that day and the dependant variable for Column 4 through 6 are the ratio of dealer sales to SOMA to total dealer buys from customer that day. Trading volumes in the denominators are restricted to CUSIPs with SOMA auctions that day. Limit Usage is defined as the ratio of VaR relative to VaR limit from the day before. Standard errors are robust.

	$\Delta p = ext{Avg}$ Price in Auction – Price _{dealer sell}					
	Full Sample			Sample of Level 1 and 2 Match		
	(1)	(2)	(3)	(4)	(5)	(6)
$\log(\text{Limit Usage})_{t-1}$	0.008***			0.009***		
	(0.002)			(0.002)		
March= $0 \times \log(\text{Limit Usage})_{t-1}$		0.006^{***}			0.006^{***}	
		(0.001)			(0.001)	
March= $1 \times \log(\text{Limit Usage})_{t-1}$		0.009^{**}			0.012^{***}	
		(0.004)			(0.004)	
Bond Life $<$ 5Yrs $\times \log(\text{Limit Usage})_{t-1}$			0.001			0.001
			(0.001)			(0.001)
Bond Life $\in [5,10)$ Yrs $\times \log(\text{Limit Usage})_{t-1}$			0.007^{*}			0.007*
$\mathbf{D} = 1 1 1 1 1 1 1 1$			(0.004) 0.021^{***}			(0.004) 0.026^{***}
Bond Life ≥ 10 Yrs × log(Limit Usage) _{t-1}			(0.021) (0.008)			(0.026^{+++})
constant	-0.002	-0.002	(0.008) -0.002	-0.001	-0.001	(0.007) -0.001
constant	(0.002)	(0.002)	(0.002)	(0.001)	(0.001)	(0.001)
	· · · ·	· · · ·	× /		```	```
Bond FE	Yes	Yes	Yes	Yes	Yes	Yes
Time FE	Yes	Yes	Yes	Yes	Yes	Yes
R^2	0.053	0.053	0.054	0.070	0.070	0.072
Ν	$11,\!012$	$11,\!012$	$11,\!007$	10,244	$10,\!244$	$10,\!239$

Table 10: Dealer SOMA Offer Aggressiveness and VaR limit Usage

Table shows the effect of log(Limit Usage) on the aggressiveness of dealer's offer in SOMA auctions. Sample period is from March 16 to June 11th, 2020. Offer aggressiveness is measured as the difference between average prices of successful offers in the auction and the dealer' actual offer (dealer sell price). The higher this measure is, the less the dealer is willing to accept when selling to the Fed, the more aggressive is his offer. Column (1) through (3) use the full sample of transactions that are identified as dealer sell to the Fed (details of the matching method described in Appendix A). Column (4) through (6) use only the sample from level 1 and level 2 match, which match the transactions and SOMA auctions data with a high degree of confidence. Limit Usage is defined as the ratio of VaRrelative to VaR limit from the day before. Standard errors are robust.

	Log (Dealer Buy from Customer)					
	In SOMA (1)	Not in SOMA (2)	Total (3)	In SOMA (4)	Not in SOMA (5)	Total (6)
Log (Limit Usage) _{$t-1$}	-0.137^{**} (0.054)	-0.116^{*} (0.063)	-0.143^{***} (0.046)	$ -0.140^{***} \\ (0.054)$	-0.121^{**} (0.062)	-0.147^{***} (0.044)
Dealer Sell to SOMA	0.046^{***} (0.017)	0.031 (0.021)	0.030^{**} (0.013)		× /	· · · ·
March=0 \times Dealer Sell to SOMA		· · · ·	()	0.036^{**} (0.018)	0.013 (0.019)	0.015 (0.014)
March=1 \times Dealer Sell to SOMA				0.083*** (0.018)	0.095^{**} (0.038)	0.083^{***} (0.015)
constant	$\begin{array}{c} 14.568^{***} \\ (0.109) \end{array}$	$\begin{array}{c} 14.674^{***} \\ (0.140) \end{array}$	15.434^{***} (0.092)	$\begin{array}{c c} 14.552^{***} \\ (0.106) \end{array}$	$14.647^{***} \\ (0.127)$	15.411^{***} (0.084)
Dealer FE	Yes	Yes	Yes	Yes	Yes	Yes
Time FE	Yes	Yes	Yes	Yes	Yes	Yes
R^2	0.884	0.851	0.908	0.885	0.854	0.910
N	506	506	506	506	506	506

Table 11: Dealer Liquidity Provision to Clients and SOMA Operations

Table shows the effect of log(Limit Usage) and the SOMA operation on dealer's ability to supply liquidity to client by taking bonds off their hands. Sample period is from March 16 to June 11th, 2020. Dependant variables in Columns (1) and (4) are Log volume of trades that are dealer buy from customers for bonds that are in the list of bonds with non-zero SOMA auction volume that day. Dependant variables in Columns (2) and (5) are Log volume of trades that are dealer buy from customers for bonds with non-zero SOMA auction volume that day. Dependant variables in Columns (2) and (5) are Log volume of trades that are dealer buy from customers for bonds that are not in the list of bonds with non-zero SOMA auction volume that day. Dependant variables in Columns (3) and (6) are Log volume of trades that are dealer buy from customers for all bonds. Limit Usage is defined as the ratio of VaR relative to VaR limit from the day before. Standard errors are robust.

	(a) h=1		
	(1) Total	(2) New	(3) Existing
log(Limit Usage)	0.109^{***} (0.018)	0.074^{***} (0.014)	$0.025 \\ (0.023)$
R^2 N	$0.213 \\ 16667$	$0.153 \\ 16665$	$0.178 \\ 16665$
	(b) h=5	ò	
	(1) Total	(2) New	(3) Existing
log(Limit Usage)	$\begin{array}{c} 0.118^{***} \\ (0.015) \end{array}$	$\begin{array}{c} 0.092^{***} \\ (0.012) \end{array}$	$0.029 \\ (0.020)$
R^2 N	$0.228 \\ 16772$	$0.162 \\ 16771$	$0.180 \\ 16771$
	(c) h=10)	
	(1) Total	(2) New	(3) Existing
log(Limit Usage)	0.130^{***} (0.015)	0.102^{***} (0.012)	$0.031 \\ (0.019)$
R^2 N	$0.232 \\ 16772$	$0.174 \\ 16771$	$0.180 \\ 16771$

Table 12: Dealer Profitability and VaR constraint

Table 12 shows the effect of log(Limit Usage) on average change in profitability over the next h = 1, 5, and 10 business days. Total, new, and existing profitability is defined to be the ratio of profit and loss (PnL) from total, new, and existing positions, respectively, scaled by VaR. Total positions PnL is defined to be the sum of PnL from new and existing positions. Limit Usage is defined as the ratio of VaR relative to VaR limit from the day before. The sample period is 1/1/15-3/31/23. All regressions include desk and time fixed effects. Standard errors are robust.

	(1) OLS	$\begin{array}{c} (2) \\ 50 \text{th} \end{array}$	(3)75th	(4) 99th
MOVE	0.719^{***} (0.014)	0.646^{***} (0.020)	0.723^{***} (0.028)	1.418^{***} (0.089)
Limit Usage Index residual	0.253^{***} (0.014)	0.181^{***} (0.025)	0.235^{***} (0.031)	0.238^{***} (0.043)
R^2	0.583			
Pseudo R^2		0.274	0.378	0.502
N	2023	2023	2023	2023

(a) Yield curve fitting error

	(1) OLS	(2) 50th	(3)75th	(4) 99th
MOVE	0.737^{***} (0.015)	0.839^{***} (0.021)	0.940^{***} (0.020)	1.027^{***} (0.087)
Limit Usage Index residual	(0.085^{***}) (0.015)	(0.021) (0.091^{***}) (0.025)	(0.090^{***}) (0.016)	(0.142^{**}) (0.067)
R^2	0.552			
Pseudo R^2 N	2026	$0.347 \\ 2026$	$0.368 \\ 2026$	$0.511 \\ 2026$

(b) 10 year Treasury on the run premia

Table 13 shows results from quantile regressions for the 50th, 75th, and 99th percentiles of Treasury market illiquidity measures on interest rate volatility (MOVE) and Limit Usage Index residual (columns 2-4). OLS regression results are shown in column 1. Treasury market illiquidity is defined in Panel (a) as the average absolute nominal yield curve fit error for securities used in the curve estimation and maturing in between 2 and 10 years, excluding on-the-run and first off-the-run securities; defined in Panel (b) as the spread between the 10-year on-the-run and off-the-run Treasury yield. Limit Usage Index is a time-series average of individual desks' values of log(Limit Usage), weighted by their market share of customer transaction volumes over the past 20 business days. Limit Usage Index residual is the residual from the OLS regression of Limit Usage Index on MOVE. Quantile regression standard errors are bootstrapped. All variables are standardized in the regressions. The sample period for the regressions is 1/1/15-3/31/23.

C Figures

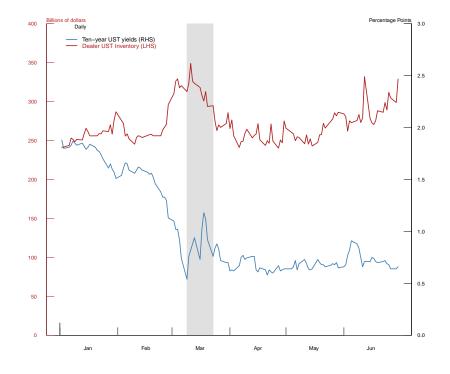
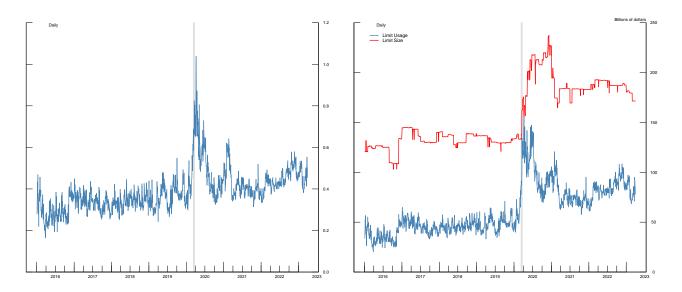


Figure 1: Dealer Inventory and 10-year Treasury Yield

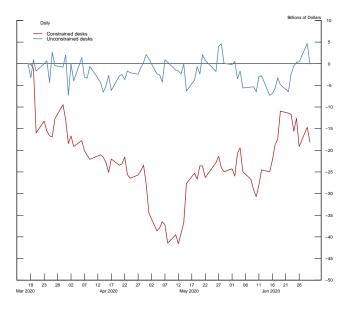
The figure plots dealer inventory of Treasury securities and the 10-year Treasury yield during the first half of 2020. Dealer inventory is measured using dealer financing of non-rehypothecated Treasury securities, including repo and other secured financing transactions. The shaded area indicates the period of Treasury market stress from March 9 to March 23, 2020. Source: FR 2052a, Bloomberg.

Figure 2: VaR Limit Size, Usage and Usage Ratio



The left figure plots the times series of average limit usage, defined as the ratio of VaR usage and the VaR limit. The right figure plots the times series of average limit size and limit usage for the VaR limit. The shaded area indicates the period of Treasury market stress from March 9 to March 23, 2020. Source: FR VV-1.

Figure 3: Inventory Change since March 16, 2020 by Proximity to VaR Constraint



The figure plots the times series of average cumulative changes since March 16, 2020 in net securities positions on UST trading desks that were relatively constrained/unconstrained by their VaR limit. More (less) constrained desks are taken as those with above (below) median VaR usage as of March 13, 2020. There are 6 desks in each group. Source: FR VV-1.

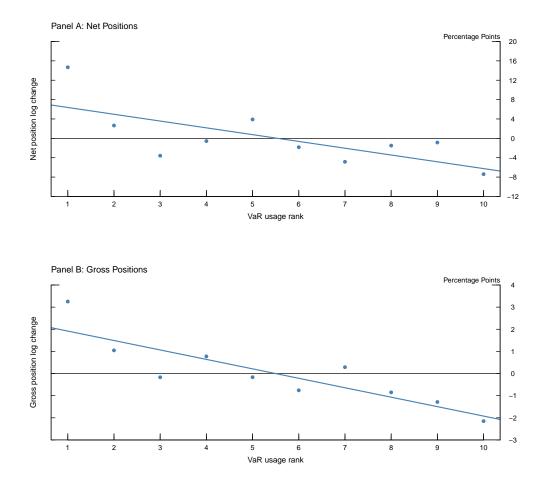
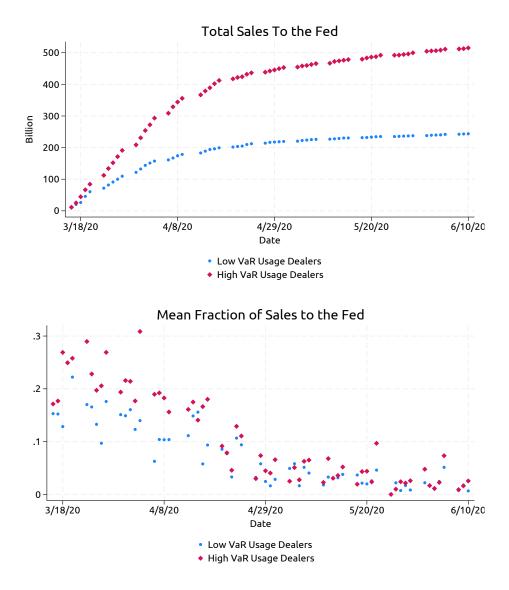


Figure 4: Position Changes against VaR Usage

The figure plots the average 10-day percentage changes in net and gross securities positions against VaR limit usage ranks. Inventory changes are averaged across decile ranks. Source: FR VV-1.

Figure 5: Sales of Treasury Securities to the Fed During Open Market Operation since March 16, 2020, Grouped By VaR limit Usage



Upper panel shows cumulative total dollar volume of dealer sales of Treasuries to the Fed during open market operations. Dealers are grouped into Low VaR usage and high VaR usage groups based on ranking of their VaR limit usages as of March 13, 2020, averaged across Treasury securities Trading desks. There are five dealers in each group. Lower panel shows the average fraction of dealer sales to clients that day that were sold to the Fed, over time.